Efficient probabilistic inversion for induced earthquake parameters in 3D heterogeneous media

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Abstract. We present an efficient probabilistic workflow for the estimation of source parameters of induced seismic events in three-dimensional heterogeneous media. Our workflow exploits a linearized variant of the Hamiltonian Monte Carlo (HMC) algorithm. Compared to traditional Markov-Chain Monte Carlo (MCMC) algorithms, HMC is highly efficient in sampling high-dimensional model spaces. Through a linearization of the forward problem around the prior mean (i.e., the "best" initial

5   model), this efficiency can be further improved. We show, however, that this linearization leads to a performance in which the output of an HMC chain strongly depends on the quality of the prior; in particular, because not all (induced) earthquake model parameters have a linear relationship with the recordings observed at the surface. To mitigate the importance of an accurate prior, we integrate the linearized HMC scheme into a workflow that (i) allows for a weak prior through linearization around various (initial) centroid locations, (ii) is able to converge to the mode containing the model with the (global) minimum misfit

10   by means of an iterative HMC approach, and (iii) uses variance reduction as a criterion to include the output of individual Markov chains in the estimation of the posterior probability. Using a three-dimensional heterogeneous subsurface model of the Groningen gas field, we simulate an induced earthquake to test our workflow. We then demonstrate the virtue of our workflow by estimating the event's centroid (three parameters), moment tensor (six parameters), and the earthquake's origin time. We find that our workflow is able to recover the posterior probability of these source parameters rather well, even when the prior

15   model information is inaccurate, imprecise, or both inaccurate and imprecise.

1   Introduction

The need to understand earthquake source mechanisms is an essential aspect in fields as diverse as global seismology (Ekström et al., 2005), oil and gas exploration (Gu et al., 2018), hazard mitigation (Pinar et al., 2003) and space exploration (Brinkman et al., 2020). In its simplest form, an earthquake source can be described, from a physics point of view, by means of a moment

20   tensor (MT) (Aki and Richards, 2002). A MT captures displacement, (potential) fault orientation, and the energy released during an earthquake. In a regional seismology context, MT inversions can provide insight into seismic afterslip patterns of megathrust earthquakes (e.g., Agurto et al., 2012). In case a seismic activity is induced by anthropogenic subsurface operations, characterizing seismic sources may also prove essential (e.g., Sen et al., 2013; Langenbruch et al., 2018). With regard to oil and gas exploration, earthquake source mechanisms are often monitored when hydrocarbons are extracted or when fluids are

25   injected into the subsurface (e.g., for fracking). In fact, such monitoring can be used to assess and mitigate the risk of ongoing
     injection processes activating existing faults (Clarke et al., 2019).

     For the purpose of monitoring induced seismicity, arrays of seismometers can be installed over the exploration area. The
     waveforms recorded by these seismometers can subsequently be exploited to characterize the induced events. For example,
     the time of the first-arrival (typically the direct P-wave) is sensitive to the earthquake hypocenter and origin time. There are

30   many inversion algorithms that exploit first arrivals to obtain estimates of earthquake hypocenters and origin times, such as the
     double-difference (Waldhauser and Ellsworth, 2000) and equal differential time (EDT) (Lomax, 2005) algorithms. However, to
     estimate MTs, it is insufficient to use only (first-arrival) travel times. In this study we therefore develop a workflow that utilizes
     full waveforms as input. Moreover, we pair the workflow with a probabilistic inversion algorithm. Consequently, the output of
     our workflow does not consist of merely best model parameters, but the so-called posterior distribution (see, e.g., Tarantola,

35   2006) of the desired earthquake parameters. We test the proposed workflow using synthetic recordings of a synthetic induced
     event in the Groningen subsurface.

     The Groningen gas field is one of the largest gas reservoirs in Europe. Since production began in 1963, more than 2115
     billion cubic meters of natural gas have been produced from the field (van Thienen-Visser and Breunese, 2015). Due to this
     gas production, the reservoir layer has compacted over time, causing earthquakes that have in some cases caused damage to

40   buildings in the Groningen province (Van Eck et al., 2006) and led to several protests against further gas extraction in the area
     (Verdoes and Boin, 2021). To investigate these earthquakes, an extensive seismometer array was installed, which is operated
     by the KNMI (The Royal Netherlands Meteorological Institute) on behalf of Nederlandse Aardolie Maatschappij (NAM)
     (Ntinalexis et al., 2019). Recently, the Dutch Research Council (NWO) established a scientific program called DeepNL to
     specifically focus on tackling the scientific challenges associated with the Groningen gas extraction. One of these challenges is

45   a thorough characterization of the source mechanism of the earthquakes. In a follow-up paper, we will validate the workflow
     proposed in this paper using field recordings of induced seismic events in Groningen.

     Event recordings collected over the Groningen field have been used as input for several inversion algorithms. Spetzler and
     Dost (2017) used the EDT algorithm to invert for the hypocenters of many Groningen earthquakes. They inverted arrival times
     of 87 events and found that all earthquakes occurred within a depth interval of 2300 to 3500 m, with most of the events

50   originating from the reservoir layer (approximately 3000 m depth). These findings are in line with the results of Smith et al.
     (2020), who used the envelopes of the seismic arrivals as input to their probabilistic algorithm. To invert for both hypocenter
     (or centroid) and MT, Willacy et al. (2018) took a different approach. Contrary to Spetzler and Dost (2017) who use a 1D
     model to represent Groningen's subsurface, they utilized a 3D heterogeneous model similar to Smith et al. (2020), and used
     the model to generate synthetic waveforms to perform a full-waveform deterministic MT inversion. The results of Willacy

55   et al. (2018), however, only focused on pure double-couple sources, which might not capture the true source dynamics. In fact,
     Dost et al. (2020) recently followed a probabilistic approach to invert event centroids and MTs of a selected number of events
     and consistently found the (non double-couple) isotropic component of the MT to be dominant and negative. The latter is in
     agreement with expectations for a compacting medium. Similar to Willacy et al. (2018), they invert waveforms but employ 1D
     local subsurface models to generate the modeled seismograms.

60   In terms of computational efficiency, each combination of a specific inversion algorithm and a specific subsurface model has both its advantages and disadvantages. In general, the main advantage of using probabilistic approaches is that the output does not consist of the best (source) model parameters but the posterior distribution of the desired earthquake parameters. Probabilistic approaches, however, are computationally expensive compared to deterministic ones. One way to reduce the computational expense is using 1D subsurface models instead of 3D velocity models to model seismograms. Unfortunately, this

65   can adversely affect the reliability of the obtained posterior because part of the heterogeneity of the subsurface is not accounted for (Hingee et al., 2011; Hejrani et al., 2017). In our workflow, we therefore deploy a computationally efficient probabilistic algorithm to invert for centroid (three coordinate components), origin time, and MT (six independent MT components) while at the same time utilizing a detailed 3D subsurface model.

      The algorithm used in our workflow is the Hamiltonian Monte Carlo (HMC) algorithm, which, for sampling high-dimensional

70   posterior distributions, has shown to be significantly more efficient than the conventional probabilistic Metropolis-Hasting algorithm (Betancourt, 2017). Using relatively low frequencies (below 0.1 Hz) and available prior information, Fichtner and Simutė (2018) developed a variant of the HMC and demonstrated its efficiency to invert for the source parameters of a tectonic earthquake. In contrast to tectonic earthquakes, where prior information regarding the event's MT, centroid and origin time is often available, such prior information is usually absent for induced earthquakes. In addition, the frequency content of induced

75   earthquakes waveforms is usually significantly higher compared to waveforms from tectonic events (Dais et al., 2018). Since higher frequencies imply shorter wavelengths, an increase in non-linearity between the observed seismograms and non-linear source parameters such as centroid and origin time can be expected in the case of induced seismic events.

      The two challenges put forward above (weaker prior and increase in non-linearity) are met by means of a workflow in which the initial prior is updated before running the HMC algorithm. In addition, multiple chains of the HMC variant are run

80   sequentially, where the results of the current chain serve as priors for the next chain. This iterative HMC is meant to provide improved prior information resulting in an adequate linear approximation. We demonstrate the validity of our workflow using data from a synthetically generated induced earthquake, which was simulated using the velocity model of the Groningen subsurface.

      In what follows, we first introduce the forward problem of obtaining surface displacements (recorded wavefields) due to

85   induced seismic source activity, including the description of a seismic source in terms of elementary moment tensors. Subsequently, we introduce the Bayesian formulation and detail the linearized HMC algorithm. Afterward, we proceed with the description and implementation of our workflow, which involves several steps that are specific to the characterization of induced seismic sources. We then test the proposed workflow using synthetic recordings of an induced earthquake source. We end by giving a perspective discussion of our results, including an outlook of applying our workflow to actual field recordings

90   of induced earthquakes from the Groningen gas field.

2 Forward problem

As with all Markov Chain Monte Carlo algorithms, HMC involves an evaluation of forward modeled data against observed data. In our case, this evaluation is between (forward) modeled surface displacement and observed displacement. Specifically, we compute synthetic displacement seismograms under the assumption of a point source. We start by writing the surface displacement $\mathbf{u}$ recorded along axial component i, due to a time-dependent force f acting in the j direction as (Shearer, 2019; Hejrani et al., 2017)

$$u_i(\mathbf{x}_r, t) = G_{ij}(\mathbf{x}_r; \mathbf{x}_a, t) * f_j(\mathbf{x}_a, t), \tag{1}$$

where $\mathbf{x}_r$ is the location at which $\mathbf{u}$ is recorded, $\mathbf{x}_a$ the source location (i.e., the location of the force f_j), and $*$ represents temporal convolution. $\mathbf{G}$ is the Green's function, and its first subscript represents its recorded component, while the second subscript indicates the direction in which the impulsive (delta function) force is acting. In the context of this study, $\mathbf{x}_r$ is the location of a KNMI seismometer.

Instead of repeatedly computing $\mathbf{u}$ for each source-receiver pair location, it is convenient to exploit source-receiver reciprocity. That is, we exploit the fact that $G_{ij}(\mathbf{x}_r; \mathbf{x}_a, t) = G_{ji}(\mathbf{x}_a; \mathbf{x}_r, t)$ (Aki and Richards, 2002; Wapenaar and Fokkema, 2006), which yields,

$$u_i(\mathbf{x}_r, t) = G_{ji}(\mathbf{x}_a; \mathbf{x}_r, t) * f_j(\mathbf{x}_a, t). \tag{2}$$

In our case, source-receiver reciprocity is numerically implemented by having each (stationary) receiver act as a source while all points of the predetermined subsurface grid covering the source region act as receivers. In addition, the behavior of the source (represented by $\mathbf{f}(\mathbf{x}_a, t)$), is in practice included in the computation of $\mathbf{G}(\mathbf{x}_a; \mathbf{x}_r, t)$ (e.g., similar to Hejrani et al. (2017)).

In case that the displacement is due to a couple of body forces instead of a single body force, displacement can be written as a sum of the displacement resulting from single forces $f_j(\mathbf{x}_a + d\hat{\mathbf{k}}, t)$ and $f_j(\mathbf{x}_a - d\hat{\mathbf{k}}, t)$ (Vavryčuk, 2013), i.e., the sum of $G_{ij}\left(\mathbf{x}_r, t; \mathbf{x}_a + d\hat{\mathbf{k}}\right) * f_j(\mathbf{x}_a + d\hat{\mathbf{k}}, t)$ and $G_{ij}\left(\mathbf{x}_r; \mathbf{x}_a - d\hat{\mathbf{k}}, t\right) * f_j(\mathbf{x}_a - d\hat{\mathbf{k}}, t)$, respectively. Here, d is the distance away from $\mathbf{x}_a$ (i.e., the force arm), and $\hat{\mathbf{k}}$ is a unit vector pointing in the k direction (k, similar to i and j, taking on values 1, 2, and 3). Furthermore, since linear momentum has to be conserved (induced seismic rupturing is an internal process) $f_j(\mathbf{x}_a + d\hat{\mathbf{k}}, t) = -f_j(\mathbf{x}_a - d\hat{\mathbf{k}}, t)$. Considering subsequently the limiting case of $d \to 0$,

$$u_i(\mathbf{x}_r, t) = \lim_{d \to 0} \left(f_j(\mathbf{x}_a + d\hat{\mathbf{k}}, t) * \left(G_{ij}(\mathbf{x}_r; \mathbf{x}_a + d\hat{\mathbf{k}}, t) - G_{ij}(\mathbf{x}_r; \mathbf{x}_a - d\hat{\mathbf{k}}, t) \right) \right) \tag{3}$$

$$= M_{jk}(\mathbf{x}_a, t) * G_{ij,k}(\mathbf{x}_r; \mathbf{x}_a, t).$$

where

$$G_{ij,k}(\mathbf{x}_r; \mathbf{x}_a, t) \equiv \lim_{d \to 0} \frac{G_{ij}\left(\mathbf{x}_r; \mathbf{x}_a + d\hat{\mathbf{k}}, t\right) - G_{ij}\left(\mathbf{x}_r; \mathbf{x}_a - d\hat{\mathbf{k}}, t\right)}{2d}, \tag{4}$$

$$M_{jk}(\mathbf{x}_a, t) \equiv \lim_{d \to 0} 2d_k f_j(\mathbf{x}_a, t), \tag{5}$$

and where d_k indicates that the infinitesimal distance d between $\mathbf{x}_a$ and $\mathbf{x}_a + d\hat{\mathbf{k}}$ is measured in the k direction. The factor two in equation (5) stems from the fact that each element M_{jk} is the result of a couple of opposite forces in the j-direction (infinitesimally separated by $2d_k$)

Together, the 9 constants M_{jk} constitute the second order seismic moment tensor $\mathbf{M}$. The MT effectively approximates a
125   seismic source by collapsing it into a single point. Due to conservation of angular momentum, The MT has only six independent components (e.g., Aki and Richards, 2002; Jost and Herrmann, 1989). Using equation (3), we can write the displacement $u_i(\mathbf{x}_r, t)$ due to $\mathbf{M}$ as follows

$$u_i(\mathbf{x}_r, t) = \sum_{j,k} G_{ij,k}(\mathbf{x}_r; \mathbf{x}_a, t) * M_{jk}(t) = \sum_{j,k} G_{ji,k}(\mathbf{x}_a; \mathbf{x}_r, t) * M_{jk}(t), \tag{6}$$

where source-receiver reciprocity has been applied to obtain the second equality. It should be understood that the spatial
130   derivatives act on $\mathbf{x}_a$, i.e., on the source coordinate of $G_{ij,k}(\mathbf{x}_r; \mathbf{x}_a, t)$ and the receiver coordinate of $G_{ji,k}(\mathbf{x}_a; \mathbf{x}_r, t)$.

To facilitate the computation of seismograms for a specific $\mathbf{M}$, we follow the work of Mustać and Tkalčić (2016) who use six independent tensors that they call elementary moment tensors as decomposed by Kikuchi and Kanamori (1991)

$$\mathbf{M}^1 = \begin{pmatrix} 0 & 1 & 0 \\ 1 & 0 & 0 \\ 0 & 0 & 0 \end{pmatrix} \qquad \mathbf{M}^2 = \begin{pmatrix} 0 & 0 & 1 \\ 0 & 0 & 0 \\ 1 & 0 & 0 \end{pmatrix} \qquad \mathbf{M}^3 = \begin{pmatrix} 0 & 0 & 0 \\ 0 & 0 & -1 \\ 0 & -1 & 0 \end{pmatrix}$$

$$. \tag{7}$$

$$\mathbf{M}^4 = \begin{pmatrix} 0 & 0 & 0 \\ 0 & 0 & -1 \\ 0 & -1 & 0 \end{pmatrix} \qquad \mathbf{M}^5 = \begin{pmatrix} -1 & 0 & 0 \\ 0 & 0 & 0 \\ 0 & 0 & 1 \end{pmatrix} \qquad \mathbf{M}^6 = \begin{pmatrix} 1 & 0 & 0 \\ 0 & 1 & 0 \\ 0 & 0 & 1 \end{pmatrix}$$

Under the assumption that each of these elementary moment tensors have the same time dependence, a specific M can be
135   described as a linear combination of these elementary moment tensors, i.e.,

$$\mathbf{M} = \sum_{n=1}^{6} a_n \mathbf{M}^n = \begin{pmatrix} -a_4 + a_6 & a_1 & a_2 \\ a_1 & -a_5 + a_6 & -a_3 \\ a_2 & -a_3 & a_4 + a_5 + a_6 \end{pmatrix}. \tag{8}$$

where the coefficients a_n ($n = 1, 2, ..., 6$) are usually referred to as expansion coefficients. In this study, we assume instantaneous rupturing of the source. This is not an uncommon assumption for (relatively small) induced seismic events. This assumption implies that the time dependence of a MT is modeled using a Heaviside function. Using the decomposition above
140   and source-receiver reciprocity, we compute elementary seismograms S_i^n

$$S_i^n(\mathbf{x}_r; \mathbf{x}_a, t) = \sum_{j,k} G_{ji,k}(\mathbf{x}_a; \mathbf{x}_r) * M_{jk}^n(t). \tag{9}$$

Consequently, we obtain

$$u_i(\mathbf{x}_r; \mathbf{x}_a, t) = \sum_{n=1}^{6} a_n S_i^n(\mathbf{x}_r; \mathbf{x}_a, t). \tag{10}$$

In practice, all $S_i^n(\mathbf{x}_r; \mathbf{x}_a, t)$ are computed for a finite number of $\mathbf{x}_a$ on a predetermined subsurface grid, separated by a distance
145 d. We detail the numerical implementation of computing the $S_i^n(\mathbf{x}_r; \mathbf{x}_a, t)$ further below.

3 Hamiltonian Monte Carlo

The HMC algorithm originated from the field of classical mechanics and its application to statistical mechanics (Betancourt,
2017). Over time, the HMC algorithm has become one of the most efficient probabilistic algorithms within the Markov Chain
Monte Carlo (MCMC) family. For our workflow, we apply a variant of the HMC algorithm that utilizes a linearization of
150 the forward problem. Therefore, we include several initial steps in our workflow to obtain priors that enable meaningful
linearization. In total, our workflow estimates ten source parameters. These are the centroid $\mathbf{x}_a$ (three components), the origin
time T0, and the MT (6 independent MT components).

Similar to other probabilistic algorithms, HMC is deployed in the context of Bayesian inference. The objective of Bayesian
inference is to obtain an estimate of the posterior probability distribution $\rho(\mathbf{m}|\mathbf{d})$ that approaches the true posterior probability
155 distribution (from here on, we will refer to $\rho(\mathbf{m}|\mathbf{d})$ as being 'the posterior'). This approach combines the likelihood $\rho(\mathbf{d}|\mathbf{m})$ of
the observed data given the modeled data with the simultaneous assimilation of the distribution of prior knowledge $\rho(\mathbf{m})$, i.e.,

$$\rho(\mathbf{m}|\mathbf{d}) \propto \rho(\mathbf{d}|\mathbf{m})\rho(\mathbf{m}), \tag{11}$$

where $\mathbf{m}$ is a vector that contains model parameters and $\mathbf{d}$ a vector containing the observed data. The likelihood evaluates a
model $\mathbf{m}$ against the observed data $\mathbf{d}$ by evaluating the misfit between the latter and forward modeled data associated with $\mathbf{m}$.
160 The HMC algorithm relies on the recursive calculation of two quantities. These are the potential energy U, which explic-
itly quantifies $\rho(\mathbf{m}|\mathbf{d})$, and the kinetic energy K, which is a function of momentum vector $\mathbf{p}$. Together, they make up the
Hamiltonian $H(\mathbf{m}, \mathbf{p})$, which represents the total energy of a system (Neal et al., 2011), and is written as follows:

$$H = K(\mathbf{p}) + U(\mathbf{m}), \quad \text{with} \quad U(\mathbf{m}) = -\ln\rho(\mathbf{m}\,|\,\mathbf{d}) \quad \text{and} \quad K(\mathbf{p}) = \mathbf{p}^T \mathcal{M}^{-1} \mathbf{p}/2. \tag{12}$$

A model $\mathbf{m}$ can be interpreted as the position of a particle within *phase space*. The phase space has a dimension that is
165 twice the dimension of the model space (i.e., this dimension coincides with the length of the vector $\mathbf{m}$ multiplied by two). By
having the same dimension as $\mathbf{m}$, the elements of the auxiliary momentum vector $\mathbf{p}$ are therefore needed to complement each
dimension of the model space (Betancourt, 2017). The movement of the particle is highly dependent on the mass matrix $\mathcal{M}$,
which therefore often acts as a tuning parameter. The mass matrix affects the "distance" a particle travels and ideally coincides
with the posterior covariance matrix. Given a certain momentum $\mathbf{p}$ the particle is allowed to travel for a certain time, τ, while
170 in conjunction fulfilling Hamilton's equations:

$$\frac{d\mathbf{m}}{d\tau} = \frac{\partial K}{\partial \mathbf{p}}, \quad \frac{d\mathbf{p}}{d\tau} = -\frac{\partial U}{\partial \mathbf{m}}. \tag{13}$$

At time τ, the particle arrives at a new location representing a new model $\mathbf{m}(\tau)$. The new model and momentum vectors are
associated with updated potential and kinetic energies, respectively, and hence a higher or lower Hamiltonian $H(\mathbf{p}(\tau), \mathbf{m}(\tau))$.

Given the probability θ, that the particle will stay at the new location, the acceptance probability is given by,

175
$$\theta = \min\left[1, \frac{\exp[-H(\mathbf{p}(\tau), \mathbf{m}(\tau))]}{\exp[-H(\mathbf{p}, \mathbf{m})]}\right]. \tag{14}$$

By recursively evaluating equations (12) to (14), we collect all locations (models) visited by the particle, except for a number of initial models (representing the burn-in period). The density of the collected models asymptotically approaches the posterior probability distribution.

In Figure 1 we exemplify the sampling behavior of both the HMC algorithm (a) and the more generic Metropolis-Hasting
180 (MH) algorithm (b) for a 2D joint probability distribution. Notice that both algorithms start with the same starting model, which is represented by the red ball. The low a posteriori probability of this initial model corresponds to a high U. The question marks in Figure 1(a) represent randomly selected models by the MH algorithm, which were not accepted due to their relatively low acceptance probability. Hence, each of these question marks involves a (computationally expensive) solution to the forward problem. Instead of using random sampling, in the HMC algorithm, the particle within phase space moves along trajectories
185 obtained by solving equation (13) leading the particle being exerted towards areas with high U, as illustrated in Figure 1(b). Furthermore, in Figure 1(b), the result of solving equation (13) which is the HMC trajectory, is represented by the brown dashed lines, and the pointing finger is the representation of momentum vector $\mathbf{p}$. For both the HMC and MH algorithms, an accepted model serves as starting model for the next sample. Although being probabilistic in terms of acceptance probabilities, the trajectories of the HMC algorithm are deterministically guided by $\partial U/\partial \mathbf{m}$ as shown in equation (13). Therefore, the algorithm
190 is also known as the Hybrid Monte Carlo algorithm (Duane et al., 1987). Thus, after proper tuning, the HMC algorithm requires less sampling than the MH algorithm to converge, which makes the HMC algorithm more computationally efficient.
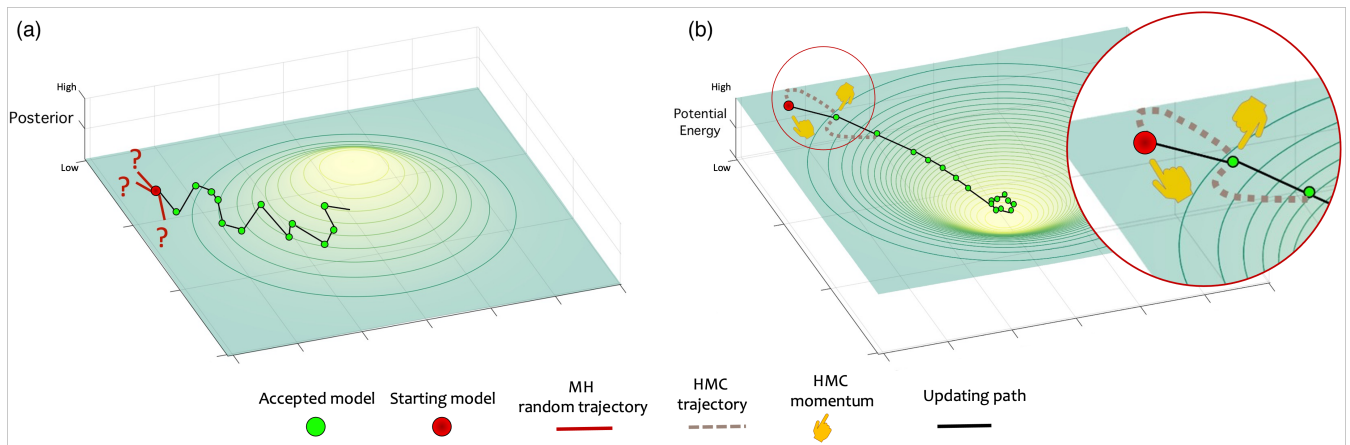
Figure 1. *Comparison between the sampling strategy of* (a) *Metropolis-Hasting and* (b) *Hamiltonian Monte Carlo* .

Assuming Gaussian distributed errors σ_d^2 in the observed data, we can write U as (Fichtner and Simutė, 2018)

$$U(\mathbf{m}) = \frac{1}{2T} \sum_{r=1}^{N_r} \sum_{i=1}^{3} \int_{0}^{T} \sigma_d^{-2} \left[u_i(\mathbf{x}_r, t; \mathbf{m}) - u_i^{\mathrm{obs}}(\mathbf{x}_r, t)\right]^2 \mathrm{d}t + \frac{1}{2N_m}(\mathbf{m} - \mathbf{m}^0)^\mathsf{T} \mathbf{C}_m^{-1}(\mathbf{m} - \mathbf{m}^0). \tag{15}$$

In our context, the $\mathbf{x}_r$ are the locations of the N_r three-component KNMI seismometers ($r = 1, 2, 3, .., N_r$). Furthermore, T
195   is the length of observed and forward modeled seismograms in time, N_m the number of model parameters (ten in our case),
$\mathbf{m}^0$ a vector containing prior means, and $\mathbf{C}_m$ the prior covariance matrix. In application to field data, the u_i^{obs} would be field
recordings by seismometers, but in this study we restrict ourselves to a numerically simulated induced event.

4   Linearization of the forward problem

In our workflow, most of the computational burden in running HMC involves the evaluation of equation (13). This is because
200   for each $d\tau$ we have to evaluate $\partial U / \partial \mathbf{m}$. To speed up the process, we use a variant of the HMC algorithm introduced by
Fichtner and Simutė (2018), in which $\mathbf{u}(\mathbf{x}_r, t; \mathbf{m})$ is approximated by means of an expansion around the prior mean, i.e.,
around $\mathbf{m}^0$:

$$u_i\left(\mathbf{x}_r, t; \mathbf{m}\right) = u_i\left(\mathbf{x}_r, t; \mathbf{m}^0\right) + \sum_{p=1}^{N_m} \frac{\partial}{\partial m_p} u_i\left(\mathbf{x}_r, t; \mathbf{m}^0\right)\left(m_p - m_p^0\right). \tag{16}$$

Substituting this linearized expression in equation (15) gives,

$$205 \quad U(\mathbf{m}) = \frac{1}{2} \sum_{p,q=1}^{N_m}\left(m_p - m_p^0\right) A_{pq}\left(m_q - m_q^0\right) + \sum_{p=1}^{N_m} b_p\left(m_p - m_p^0\right) + \frac{1}{2}c \tag{17}$$

where A_{pq}, b_p, and c read

$$A_{pq} = \frac{1}{T\sigma_d^2} \sum_{r=1}^{N_r} \sum_{i=1}^{3} \int_0^T [\frac{\partial}{\partial m_p} u_i\left(\mathbf{x}_r, t; \mathbf{m}^0\right) \frac{\partial}{\partial m_q} u_i\left(\mathbf{x}_r, t; \mathbf{m}^0\right)] \mathrm{d}t + \frac{1}{N_m \sigma_m^2}, \tag{18}$$

$$b_p = \frac{1}{T\sigma_d^2} \sum_{r=1}^{N_r} \sum_{i=1}^{3} \int_0^T \left[u_i\left(\mathbf{x}_r, t; \mathbf{m}^0\right) - u_i^{\mathrm{obs}}\left(\mathbf{x}_r, t\right)\right] \frac{\partial}{\partial m_p} u_i\left(\mathbf{x}_r, t; \mathbf{m}^0\right) \mathrm{d}t, \tag{19}$$

210   and

$$c = \frac{1}{T\sigma_d^2} \sum_{r=1}^{N_r} \sum_{i=1}^{3} \int_0^T \left[u_i\left(\mathbf{x}_r, t; \mathbf{m}^0\right) - u_i^{\mathrm{obs}}\left(x_r, t\right)\right]^2 \mathrm{d}t. \tag{20}$$

Differentiating equation (17) with respect to m_p, we have (Fichtner and Simutė, 2018)

$$\frac{\partial U}{\partial m_p} = \sum_{q=1}^{N_m} A_{pq}\left(m_q - m_q^0\right) + b_p, \tag{21}$$

which, together with the random momentum vector, determines the HMC trajectory.

215   Because the displacement depends linearly on the moment tensor components (see equations (9) and (10)), equation (16) is
exact with respect to these parameters. The dependence on the other parameters is non-linear, and this non-linearity increases

as the frequency of the input data increases. Therefore, in the case of induced events, which usually generate higher frequencies than stronger, regional events, the non-linearity is considerably higher. Hence, to have a tolerable linearization, accurate priors are required when it comes to the centroid and origin time. Without sufficiently accurate priors, the above HMC variant will 220 struggle to sample the mode containing global minimum of the potential energy. Therefore, we propose an approach that involves an initial estimation of the prior mean in order to permit this linearization. This is detailed further below.

5 Numerical implementation

In practice, the elementary seismograms discussed in Section 2 are computed for a finite number of possible centroid locations. That is, prior to our probabilistic inversion, we generate a database of these seismograms. This database contains, for each possi- 225 ble source location $\mathbf{x}_a$ and receiver location $\mathbf{x}_r$ (r=1,...,N_r), a total of 3 x 6 = 18 elementary seismograms (three components for each of the six elementary moment tensors). In our case, each $\mathbf{x}_r$ corresponds to a (KNMI) seismometer location that recorded the induced event. The elementary seismograms are computed using the spectral element software SPECFEM3D-Cartesian (Komatitsch and Tromp, 2002) and we exploited spatial reciprocity while doing so. We use an existing detailed Groningen velocity model (Romijn, 2017) for this purpose from which we construct a regular grid of the model using gnam and PyAspect 230 Python packages that are available at https://github.com/code-cullison/gnam and https://github.com/code-cullison/pyaspect.

To confirm the successful implementation of source-receiver reciprocity, we simulate a scenario of an induced event in the Groningen gas reservoir (Figure 2). The centroid is indicated with a red star, and the receivers are depicted as white triangles. At each location, the wavefield is "recorded" at 200 meters depth by the deepest of a series of four borehole geophones (Ruigrok and Dost, 2019). The elementary seismograms computed at the location of KNMI station G094 are shown in Figure 235 3 (green), and superimposed on top (yellow) are the waveforms resulting from the application of source-receiver reciprocity. All seismograms are bandpass filtered between 1 and 3 Hz, similar to the passband used by Dost et al. (2020).

We integrate the above HMC variant into our workflow by implementing a leapfrog algorithm for evaluating equation (13). Furthermore, we define $d\tau$ as suggested by Neal et al. (2011) to ensure numerical stability and set a fixed value for τ for all chains. The construction of the mass matrix $\mathcal{M}$ is discussed in the next section.

240 6 An iterative approach

The performance of the linearized HMC variant strongly depends on the prior means (see equation (16)). For that reason, we propose a workflow in which the algorithm is run iteratively, with each iteration involving an update of the priors to allow for an updated linearization. Specifically, instead of evaluating equation (18) - (20) once, we run a sequence of HMC chains. For each successive chain, the posterior means and standard deviations from the previous chain act as prior means and entries 245 for $\mathcal{M}$ in the new chain (i.e., the next iteration). For the first chain in the sequence, the "initial" prior means (i.e., $\mathbf{m}^0$) are obtained via a specific scheme integrated into the workflow. The estimation of these prior means is described in more detail in the subsection below.
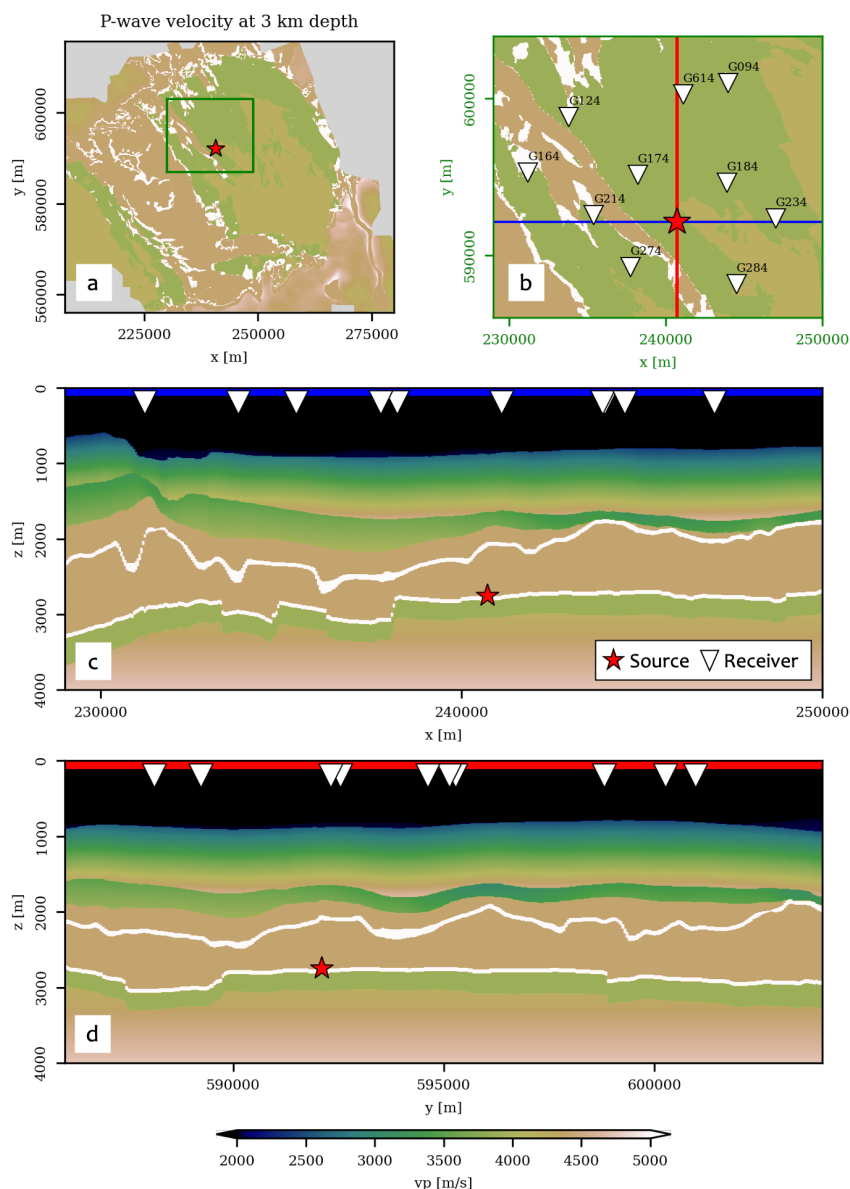
Figure 2. *Scenario of an induced earthquake in Groningen area. a. Horizontal slice of the Groningen P-wave velocity model. b. Zoom of the area indicated by the green rectangle in a. 'Inverted triangles indicate locations of KNMI stations (i.e., the $\mathbf{x}_r$)'. c. Vertical slice along the blue line in b. d. Vertical slice along the red line in b.*

We test our workflow for an induced event shown in figure 2. We set the MT components to 9E+13 Nm, -1E+13 Nm, -3E+13 Nm, 8E+13 Nm, 5E+13 Nm, and 4E+13 Nm for M_{11}, M_{22}, M_{33}, M_{12}, M_{13}, and M_{23}, respectively. Using the moment-magnitude relation given by Gutenberg (1956) and Kanamori (1977), this moment tensor can be shown to correspond

250
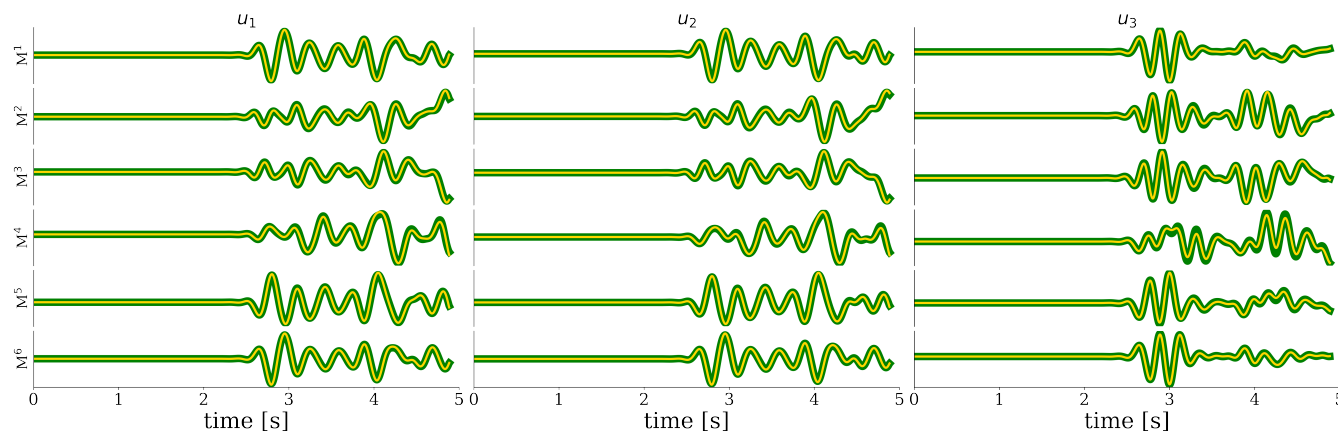
Figure 3. *Comparison between elementary seismograms due to a source at the actual location (red star) and the receiver at G094 (green) and the elementary seismograms resulting from the implementation of source-receiver reciprocity (yellow). The equality of the traces confirms successful implementation of source-receiver reciprocity. Along the vertical axis, all six (independent) elementary seismograms are depicted. Left, center, and right plots show particle displacement in the x_1, x_2, and x_3 direction, respectively.*

to an earthquake of magnitude 3.28 Mw. We also add noise to our synthetic seismograms in order to make our experiment more realistic. This noise is added in the frequency domain by multiplying the (complex) spectrum of each synthetic seismogram with a bivariate normal distribution that has a zero mean and a standard deviation of 15% of the amplitude of the seismogram at the dominant frequency. As a result, this noise will not only give amplitude variations but also varying time-shifts with
255 respect to the true synthetic seismograms. When running the Markov chains, we assume the data uncertainty (σ_d) to be 30% of the maximum amplitude of each seismogram. Admittedly, this is rather arbitrary, and in the application to field data, the data uncertainty has to be estimated from the obtained seismograms themselves. Finally, we set the origin time to 14 s.

6.1 Prior mean estimation

Before running the first HMC chain, we need to estimate the initial prior means and variances. In short, we propose an approach
260 in which a first-arrival based algorithm is used to estimate the centroid. Subsequently, the origin time can be estimated, after which equation (21) can be used to compute the prior means for the individual moment tensor components. Each of these steps is now discussed in more detail.

Numerous algorithms exist that allow one to estimate an earthquake's hypocenter and/or centroid. Here we propose using first-arrival based algorithms for this purpose since these are computationally more efficient than waveform-based algorithms.
265 First-arrival based algorithms only require the computation of the P and S-wave arrival times, and by adopting a high-frequency approximation (e.g., Aki and Richards, 2002), these arrivals can be found by running one of the various Eikonal solvers (e.g., Noble et al., 2014). For example, the EDT method detailed in Lomax (2005) can be used for this purpose (Masfara and Weemstra, 2021).

As an alternative to using a first-arrival based algorithm, the prior means of the centroid can instead be retrieved from existing

270   literature if it exists. For example, in the case of the induced seismicity in the Groningen field, Smith et al. (2020) have shown

that they could resolve hypocenters with maximum uncertainties of 150 m and 300 m for epicenter and depth, respectively.

Their results could be considered as priors. Another option is to use the epicenters from the KNMI earthquake database, which

by default has all depths set to 3 km.

Given a centroid prior mean that was either calculated or acquired from literature, the prior mean of the origin time can

275   be estimated by computing the P-wave traveltimes from a centroid prior to each of the receivers. These travel times can be

computed using the same Eikonal solver that was used to obtain the centroid prior (e.g., the Fast Marching Method; Sethian

and Popovici, 1999). By subtracting the computed travel times from the observed (picked) first-arrival times and averaging

across receivers, an initial origin time prior mean can be obtained.

To refine the initial origin time prior estimate, we cross-correlate the envelope of the observed seismograms $env[u_i^{obs}(\mathbf{x}_r, t)]$

280   with the envelope of the forward modeled seismograms $env[u(\mathbf{x}_r, t, \mathbf{m}^0)]$. We do this for each component of each receiver

location individually. The forward-modeled seismograms are computed using full-waveform modeling (detailed in Section 5)

using the initial prior means for centroid and origin time, and given arbitrary MT components. Specifically, we compute

$$T0_{shift} = \operatorname*{argmax}_{t} \sum_{r=1}^{N_r} \sum_{i=1}^{3} env[u_i(\mathbf{x}_r, t; \mathbf{m}^0)] \star env[u_i^{obs}(\mathbf{x}_r, t)], \tag{22}$$

where $T0_{shift}$ is the additional time shift that needs to be added to the initial origin time prior mean to obtain the refined origin

285   time prior.

We test equation (22) using the synthetic earthquake shown in Figure 2. For this test, we add 600 m to the true x_1, x_2, and x_3

centroid components, and we impose a (rather aggressive) 9 s time shift with respect to the true origin time. Note, this implies

that we did not employ the aforementioned procedure to obtain initial centroid and origin time priors because this would result

in a centroid and origin time estimate that would be too close to the true centroid and origin time – essentially rendering the use

290   of equation (22) unnecessary. In other words, we deliberately impose large deviations from the true values to show the merit of

using equation (22).

Given arbitrary MT components, we show in Figure 4 the result of applying equation (22) to vertical surface displacements.

In Figure 4(a), we depicts the envelopes of modeled seismograms given available prior means (i.e., 600 m deviation from the

true x_1, x_2, and x_3 values, and 9 s from the true $T0$). Figure 4(b) shows the noisy synthetic "observed" seismograms and

295   Figure 4(c) is the result of applying equation (22) to each of the displacement envelopes. In Figure 4(d), we show the result of

stacking all signals in Figure 4(c). The vertical blue line indicates the time at which the stack of the cross-correlated envelopes

attains its maximum value, i.e., $T0_{shift}$; the vertical red line represents the deviation of the initial origin time prior from the

true origin time (i.e., 9 seconds in this example).

Having sufficiently accurate prior means for the centroid and origin time, we then estimate the prior mean of the MT. For

300   this purpose, we keep the centroid and origin time constant but solve for the remaining six parameters (the independent MT

elements). In Section 4, we showed that because equation (17) is a quadratic function of $\mathbf{m}$, its derivative is linear in $\mathbf{m}$ (see

equation (21)). This first derivative hence coincides with zero for that model for which $U(\mathbf{m})$ attains its (global) minimum
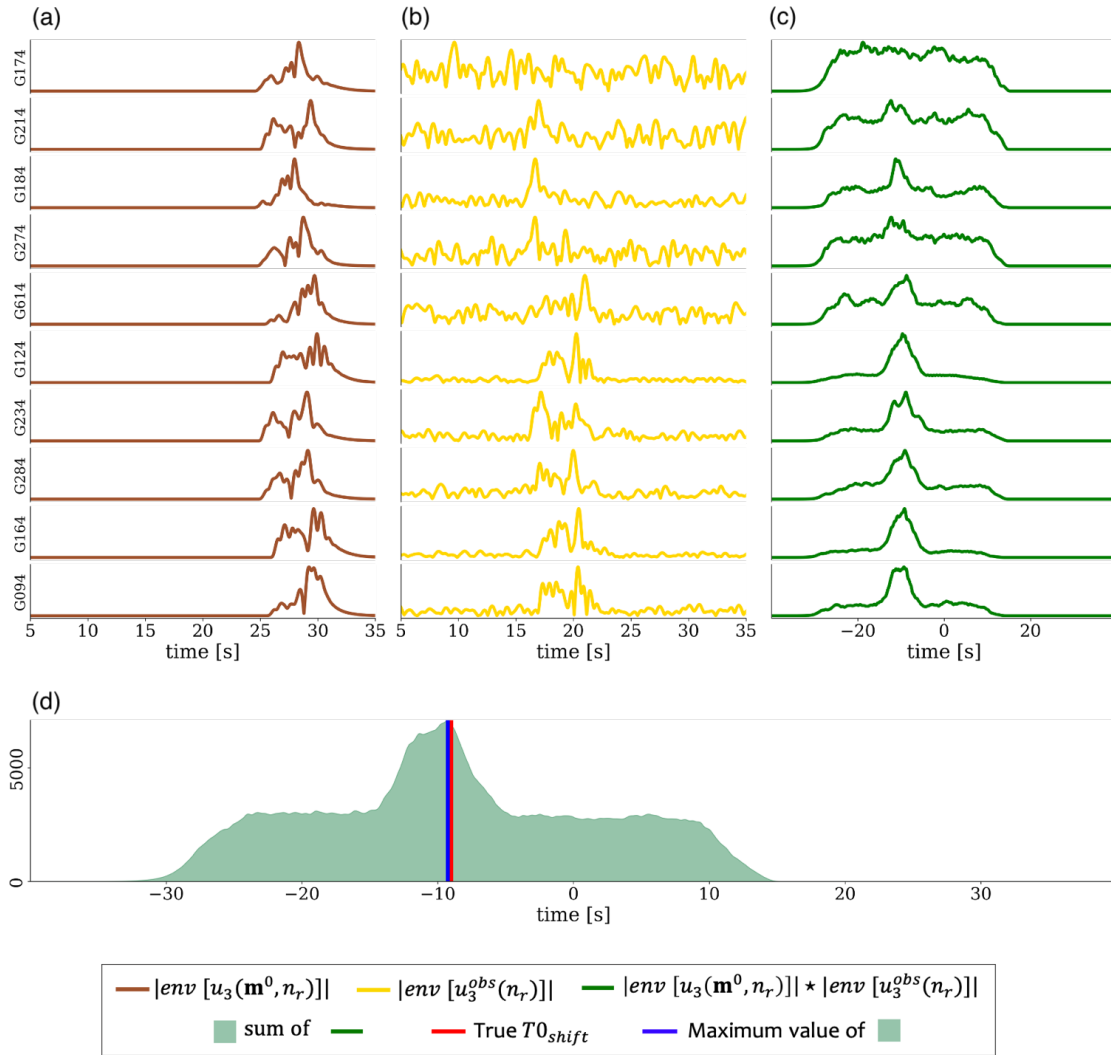
Figure 4. *The results of estimating the prior mean of origin time using equation (22) (d), given the envelopes of modeled displacements (a), noisy synthetic "observed" seismograms (b), and the convolution between (a) and (b) in (c).*

value. As such, setting this derivative to zero allows us to obtain a first estimate (i.e., prior means) of the moment tensor components. Setting the left-hand side of equation (21) to zero yields

$$\mathbf{m} = \mathbf{A}^{-1}\left(\mathbf{A}\mathbf{m}^0 - \mathbf{b}\right) = \mathbf{m}^0 - \mathbf{A}^{-1}\mathbf{b}. \tag{23}$$

305

It should be understood that equation (23) is implemented with $T0$ and centroid fixed. Hence, the model vector has only six elements, and $\mathbf{A}$ is a six by six matrix. The quadratic nature of U in equation (17) furthermore implies that arbitrary values can

be chosen for the initial moment tensor components in $\mathbf{m}^0$. In fact, in the absence of noise and the correct prior means for the centroid and origin time, the MT priors estimated using equation (23) will coincide with the true MT components.

310   In practice, the prior means resulting from equation (23) may still deviate significantly from the true values due to the inaccuracy of the initial centroid and origin time priors. Solving equation (23) nevertheless provides sufficiently accurate prior information regarding the magnitude of the induced event. Finally, it is useful to draw a parallel with typical least-squares optimization problems (e.g., Virieux and Operto, 2009). In such context, $\mathbf{A}$ is analogous to the Hessian and the difference between $\mathbf{m}$ and $\mathbf{m}^0$ in equation (23) can be considered the model update vector.

6.2   Full workflow

In Figure 5, we illustrate our entire workflow. The main component of the workflow is the iterative HMC procedure, which is preceded by the (just-described) determination of the initial prior and succeeded by the evaluation of the posteriors.
The determination of the initial prior consist of the following four steps:

1. Estimate the initial prior mean for the centroid, either by running a first-arrival based probabilistic inversion algorithm
320   or by extracting it from existing literature.

2. Estimate the initial prior mean of the origin time using (P-wave) travel times from the centroid obtained in step 1 to the receiver locations. This estimate is refined by evaluating equation (22) using an arbitrary MT.

3. Estimate the initial prior mean of the MT by fixing centroid and origin time to their prior means (steps 1 & 2) and solving equation (23). The sought-after MT prior means are contained in $\mathbf{m}$ upon substitution of arbitrary MT components in
325       $\mathbf{m}^0$.

4. Determine the standard deviation for each of the ten model parameters: centroid (3), origin time, and moment tensor (6). These standard deviations are needed to construct our first mass matrix $\mathcal{M}$. Ideally, $\mathcal{M}$ is the posterior covariance matrix. Here we approximate it by a (10x10) diagonal matrix with the following entries for the diagonal. For the first three entries (representing the centroid), we take the standard deviation of the centroid prior mean obtained in step 1.
330   For the entry representing origin time, we use half the period of the dominant frequency in the recordings. For the MT components, we use a fraction of the minimum absolute value of the MT prior means obtained by solving equation (23).

Now that the (initial) prior means and standard deviations are determined, the HMC variant is run iteratively up to n chains (yellow box in Figure 5). A test for chain convergence might be required to determine the number of chains needed, and it is highly dependent on the quality of the prior means, data uncertainty, model uncertainty, initial model, and the dominant
335 frequency of the observed recordings. In our example (detailed below), approximately ten chains are sufficient when the distance between the initial estimation of the centroid and the true centroid is less than 700 m. The separate steps of the iterative HMC procedure are:

5. Collect the prior means and associated standard deviations, and construct the mass matrix $\mathcal{M}$. For the first chain, the output from steps 1 to 4 is used as input. In subsequent chains, they are extracted from the posterior of the previous HMC

340   chain. In this step, also $\mathbf{A}$ (equation (18)), $\mathbf{b}$ (equation (19)) and c (equation (20)) are recomputed.

6. Run a new HMC chain with a preset number of iterations and burn-in period. Note, that for each chain, the results are stored for latter use.

7. Collect the results. The means and standard deviations will serve as input of for the next iteration (see step 5).

After a total of n HMC chains, we evaluate the posteriors (dark blue box in Figure 5). This involves:

345   8. For each of the n posteriors, compute the means $\mathbf{m}^s$ $(s = 1, ..., n)$. We use these means to generate synthetic recordings and evaluate them against the observed data through determination of the variance reduction (VR)

$$VR = 1 - \sqrt{\frac{\sum_{r=1}^{N_r} \sum_{i=1}^{3} \left(u_i(\mathbf{x}_r, t; \mathbf{m}^s) - u_i^{\mathrm{obs}}(\mathbf{x}_r, t) \right)^2}{\sum_{r=1}^{N_r} \sum_{i=1}^{3} \left(u_i^{\mathrm{obs}}(\mathbf{x}_r, t) \right)^2}} \tag{24}$$

9. Define a VR threshold. Posteriors associated with a $\mathbf{m}^s$ for which the VR exceeds this threshold are used to compute the final posterior distribution.
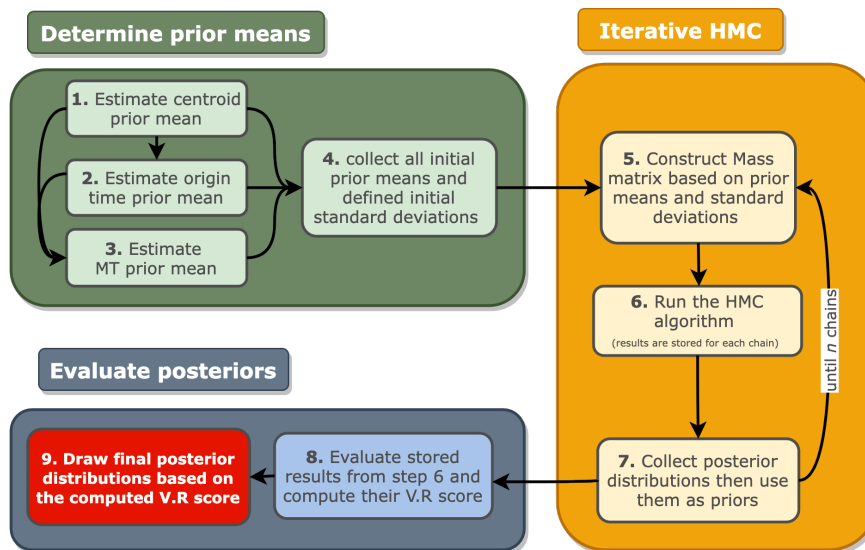
Figure 5. *Full workflow of our iterative HMC scheme.*

350   We use the above workflow to estimate the parameters of the synthetic event shown in Figure 2. In step 1, we assume a suitable prior of the centroid can be retrieved from literature (e.g., from Smith et al. (2020) in case of the seismicity in Groningen). To simulate the fact that this prior may well deviate from the true centroid, we shift this initial centroid prior mean

by 600 m in all directions (i.e., with respect to the correct event location). Having the prior mean for the centroid, we follow

steps 2 and 3 in the workflow to obtain the other prior means. To encode for a state of ignorance, we set the standard deviation

355   σ_m of each model parameter to infinity, which implies that the last term of equation (18) evaluates to 0. The elements of our

initial mass matrix are taken from the results of steps 1-3 as explained in the full workflow (step 4), except for those elements

that correspond to the centroid; these we set to 300 m. Using the initial prior means and the initial mass matrix, we run 20

chains of the HMC variant. Furthermore, we run 2500 iterations (step 6) for every chain, with the first 500 samples discarded

as burn-in samples. After finishing all iterations, the results of each current chain are then used to update the prior means and

360   mass matrix for the next HMC chain (the actual iterative HMC procedure). For each of the ten model parameters, all 40,000

samples (20 iterations × 200 samples) are depicted in Figure 6. To obtain our final posterior, we take the results of chains for

which the means are associated with seismograms yielding a VR $>= 85\%$ of the maximum VR. In detail, the samples from

all chains (black dots) and the selected chains (green dots) are depicted in Figure 7. For the samples of the selected chains, the

one-dimensional marginal probability distributions of each of the ten source parameters are shown in Figure 8. In Figure 9,

365   we show the synthetic (observed) seismograms (green), seismograms generated from the final posterior means (blue), and the

true noise-free seismograms (red). With the use of a database containing pre-computed elementary seismograms and using the

python code we developed, the entire workflow takes approximately one minute to finish on a single-core CPU system.
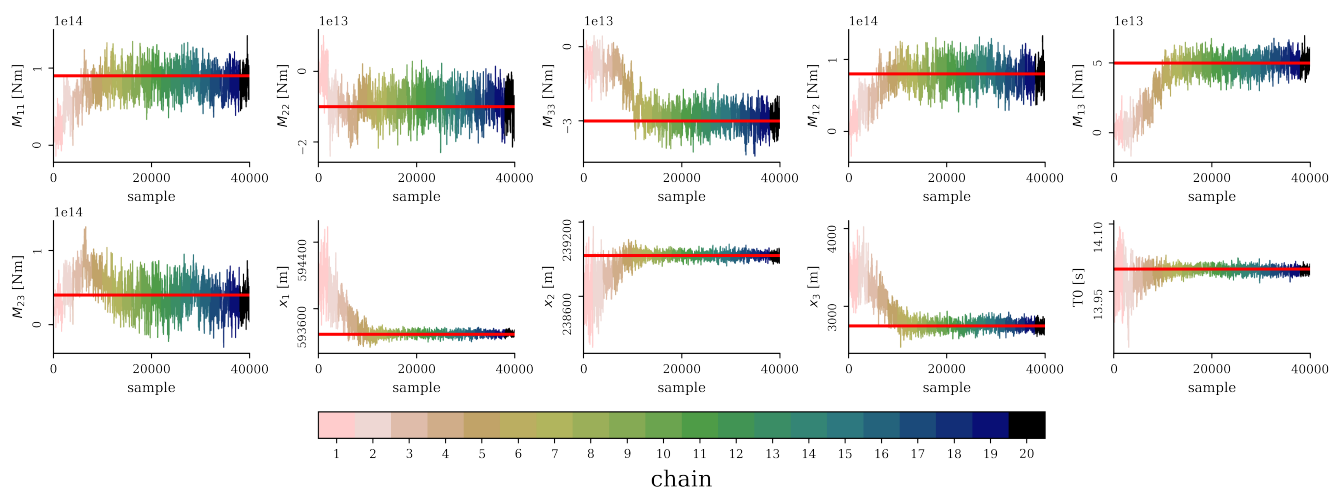
Figure 6. *Results of our iterative HMC scheme for a total of 20 chains, each involving 2500 steps of which the first 500 are discarded as burn-in samples (not shown). The red lines are the true values.*

7   The importance of the prior

The above workflow might not be optimal if the initial prior information is "weak" in the sense that the initial centroid prior

370   mean deviates significantly from the true value. This is due to the fact that our forward problem is in essence a non-linear
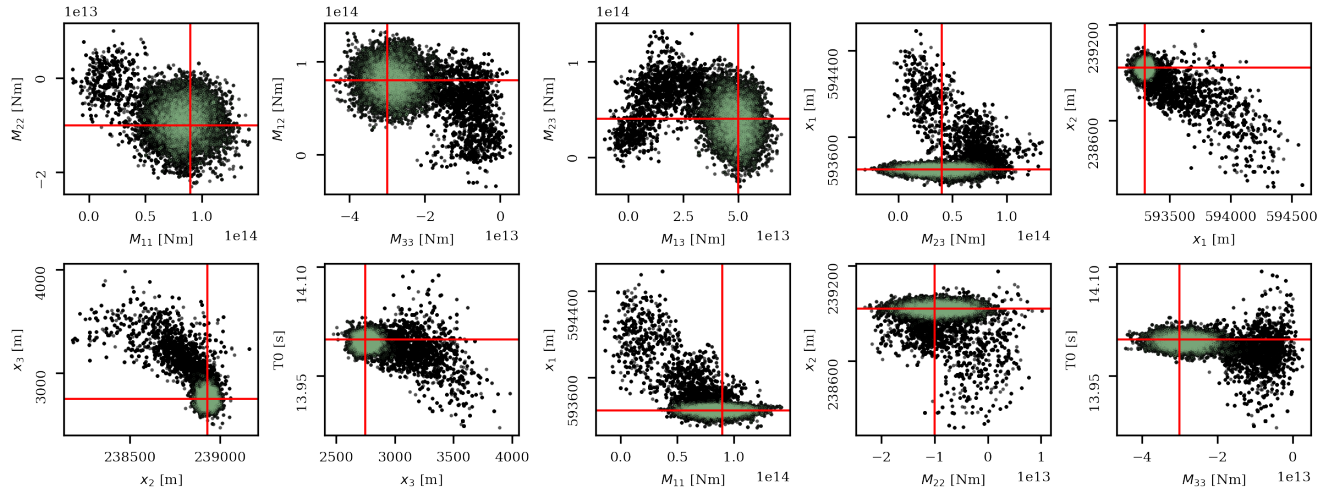
Figure 7. *Ten two-dimensional marginal probability densities of the inverted model parameters. Black dots are all the samples given the results from all chains (Figure 6), whereas the green dots represent the samples from chains that give VR>= 85% of the maximum VR, which then represent our final posterior. The red lines are the true values.*
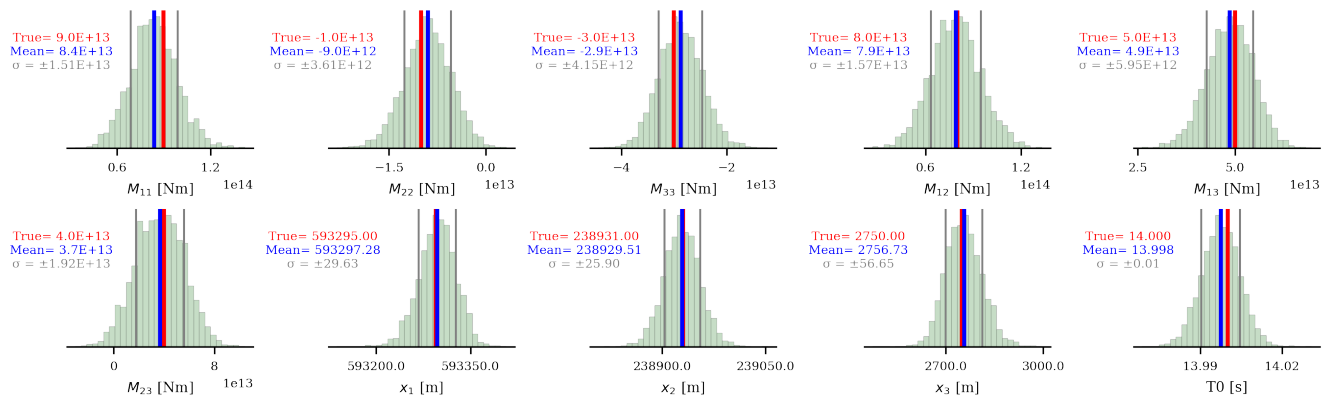
Figure 8. *The final marginal posterior distributions (green samples in Figure 7). The means are represented by the blue lines and the gray lines are the standard deviations. Red lines are the true values.*

problem, whereas the adopted linearization (Section 4) relies on the assumption of it being only weakly non-linear. In other words, poor initial centroid priors imply linearization around a location $\mathbf{x}$ that deviates too much from the true source location $\mathbf{x}_a$, which may result in the HMC algorithm "getting stuck" in local minima. This problem can be mitigated by running the workflow with multiple initial prior means. Depending on how close each of the initial prior means is to the true values, some

375 chains might get stuck in a local (minimum) mode while others correctly sample the mode containing the global minimum (or
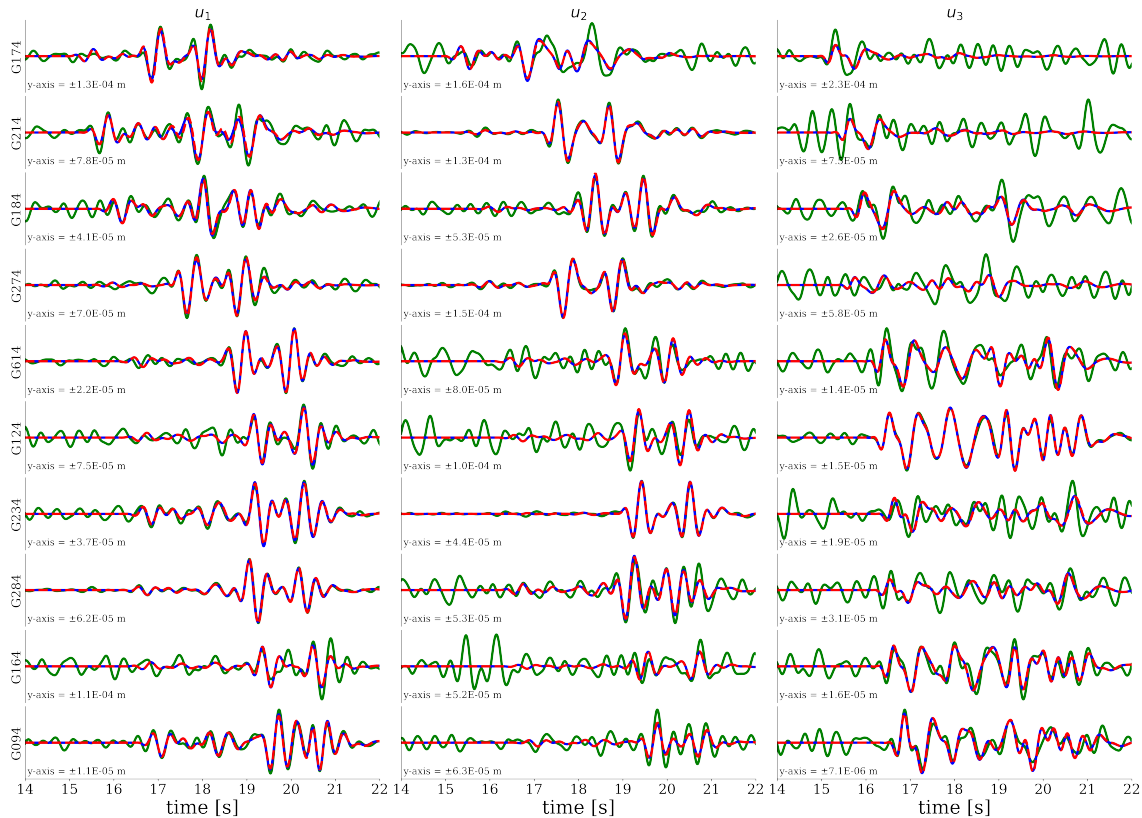
Figure 9. *Comparison between the true seismograms (red), observed (true+noise) seismograms (green), and seismograms generated using the final posterior mean (blue).*

global maximum if one considers $\rho(\mathbf{m}|\mathbf{d})$. In the end, the final posterior can be drawn by combining the results of all chains given multiple initial prior means.

To showcase the effect of weak prior information in the context of induced seismicity in the Groningen gas field, we re-use the synthetic earthquake in Figure 2. However, instead of shifting the initial centroid prior mean by 600 m for all coordinate
380   components, we rigorously shift it by 1 km for each horizontal coordinate. Meanwhile, the depth is set to 3 km, corresponding to the default depth in the KNMI database. In particular because, in application to field data, this database will be our primary source to obtain our priors. To get additional initial centroid prior means, we construct a 2.8 x 2.8 km 2D grid, at a depth of 3 km, with a spacing of 700 m, centered around the initial centroid prior mean. A pre-test can help in determining the grid spacing. We previously demonstrated that our workflow performs well when the centroid prior means are shifted by 600 m (in
385   all directions) from their true values. This shift corresponds to an absolute deviation of about 1 km. Given the spacing of the constructed grid, assuming that the depth could be around $\pm$ 500 m, the maximum total distance is about 700 m, which is then considered safe for the HMC algorithm to sample the mode containing the global minimum.

Overall, given our 5x5 horizontal grid, we have 25 initial centroid prior means, each of them being subjected to our workflow. For each workflow run, we use the same model and data uncertainty as in the initial synthetic case. The same applies to the

390   number of chains (20), samples per chain (2500), and burn-in period (500 samples). To reduce the computational time, we run the 25 workflows (associated with 25 initial centroid prior mean) simultaneously by parallelizing our code. We subsequently collect the results of each workflow to obtain an estimate of our posterior distribution. The results of this parallelization are summarized in Figure 10, which highlights the effect of the separation between the centroid prior mean and the true centroid. Using the same threshold as in our initial experiment (VR$>= 85\%$ of maximum VR), in Figure 11, we show all (non burn-in)

395   samples associated with the selected chains and samples from all chains given the calculated VR. For the selected chains, the marginal probability distribution of each parameter is presented in Figure 12. As expected, chains with an initial centroid prior mean relatively close to the true centroid converge to the true mode (containing the global minimum). At the same time, chains starting from a centroid further away from the true centroid, "got stuck" in a local mode. Fortunately, our VR strategy is still successful in picking appropriate chains, allowing us to obtain an estimate of the posterior distributions.
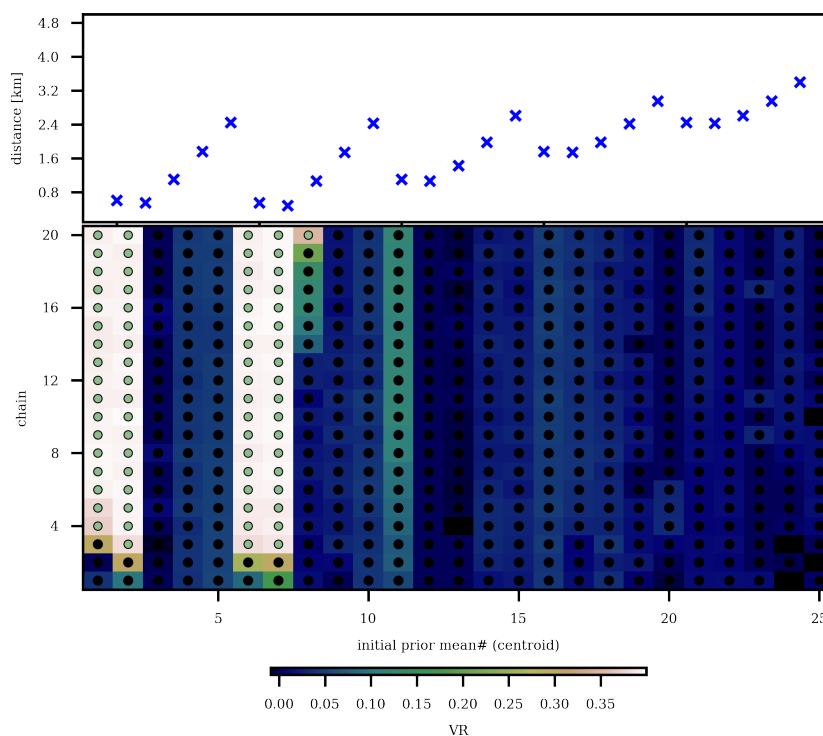
Figure 10. *Summary of running the workflow using the 25 initial prior means. The distance between each of the initial centroid prior means and the true centroid is indicated on top. Below, we show for each of these initial prior means the VR as a function of chain number (vertical axis). Chains associated with a VR$>= 85\%$ of the maximum VR (0.4) are labeled with green dots, whereas chains with a posterior mean yielding seismograms for which the VR does not exceed 85%, are labeled with black dots.*
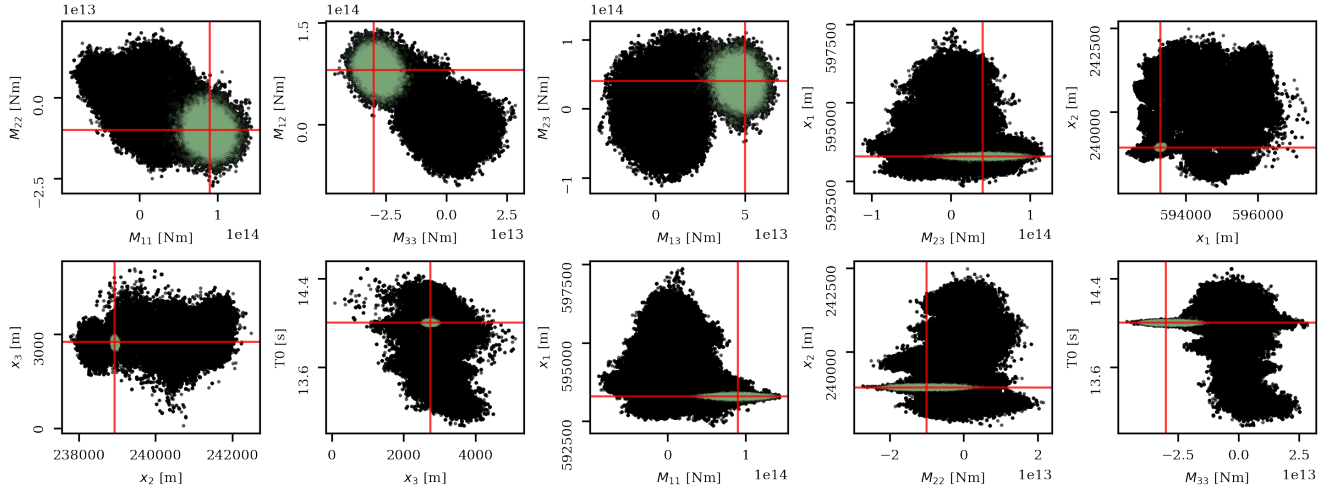
Figure 11. *Same ten two-dimensional marginal probability densities as in Figure 7). Note that scales on horizontal and vertical axes differ. Black dots are the samples from all chains, whereas the green dots represent the samples from chains with a posterior mean that yields a VR that is higher than 85% of the maximum VR. The red lines represent the true values.*
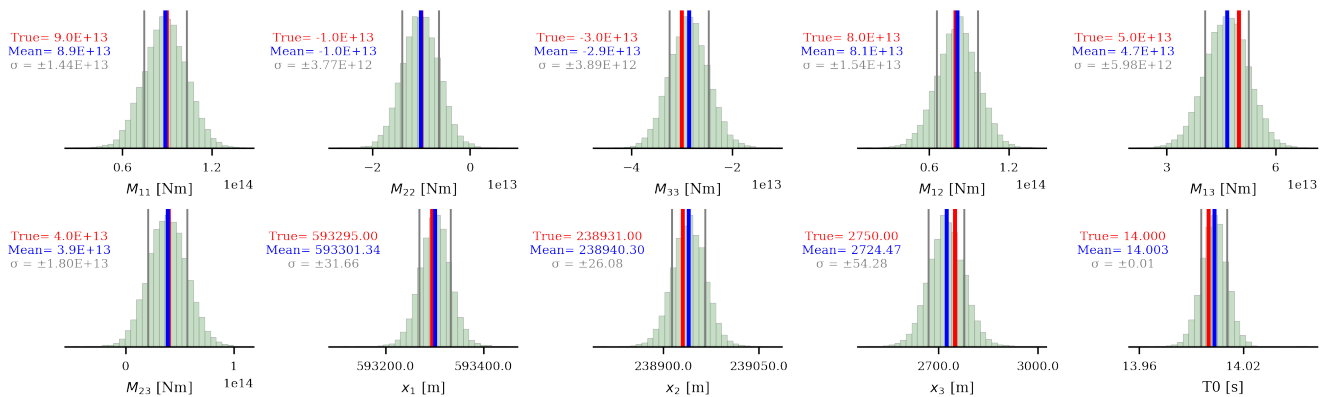
Figure 12. *Marginal posterior distributions for each model parameter given the selected chains (depicted as green dots in Figure 10). The blue lines represent the means, and the standard deviations are represented by the grey lines. The red lines represent the true values.*

8 Discussion and Conclusions

Using synthetic events, we demonstrate that our probabilistic workflow is able to estimate the posterior probability of the various parameters describing induced seismic events. Our workflow includes a systematic approach to obtain meaningful initial priors, which is particularly important for the employed HMC variant: the linearization of the forward problem around the prior mean requires the initial priors to be sufficiently close to the true event location. Furthermore, we show that by using

405 an iterative scheme, we can update the prior mean such that convergence is obtained to a centroid location that allows the estimation of a meaningful posterior. The iterative scheme involves sequentially updating the prior mean of each new HMC chain using the posterior estimate obtained from the previous HMC chain. This approach is based on the suggestion of Fichtner and Simutė (2018) to repeat the Taylor expansion (of the forward problem) for each new sample of the Markov chain. However, we (only) do this every 2500 samples in our case. A brute-force approach to perform the expansion at each step on every chain

410 would make our computational costs prohibitively large.

Prior to executing the workflow, one needs to compile a database of the elementary seismograms, which often requires significant computing power. Once compiled, our workflow can be run efficiently. Using a single-core CPU system, a single run of our workflow with 20 HMC chains takes about one minute to finish, where each chain consists of 2500 iterations. In contrast, the computational costs of the MH algorithm (to get the same results) would be much higher, as previously shown by

415 (Fichtner and Simutė, 2018; Fichtner et al., 2019). Furthermore, various modifications could be applied to the workflow, such as adding simulated annealing and tempering (Tarantola, 2006), including a step to quantify the error in the input seismograms (Mustać and Tkalčić, 2016), and/or applying a scheme that is able to tune $d\tau$ and τ for each HMC run (Hoffman et al., 2014). These modifications could be beneficial, especially when dealing with field observations, which is the subject of future work. We also show that the workflow can be adapted to account for scenarios in which the initial centroid prior mean is rather

420 inaccurate and/or the initial prior is weak. If that is the case, an approach can be adopted in which various iterative HMC workflows, each using a different centroid prior mean as a starting point, are run. Subsequently, using the variance reduction associated with the posterior means of the individual chains as a binary criterion for selecting a chain's samples, a final estimate of the posterior probability can be obtained.

We would like to emphasize that our workflow is, in principle, not limited to inversions of the parameters we use here.

425 We could extend our probabilistic inversion to parameters such as stress drop, velocity, or to invert for finite fault source parameters. Furthermore, it is important to mention that our workflow aims to invert seismic source parameters using seismic surface recordings in a specific frequency range. That is, it is specifically geared towards inverting for induced seismic events. We found that the workflow works well when applied to data with frequencies between 1 and 3 Hz. For higher frequencies, however, some testing might be needed because the non-linearity between the input data and model parameters increases with

430 frequency. Finally, it is important to note that, in this study, we have assumed the velocity model to be known and correct. This means that, in application to field data, the estimated posterior probability gives a lower bound in terms of the variability of the source parameters: inaccuracies in the velocity model necessarily imply broader posterior probabilities.

Code availability. The seismogram database was generated using spectral element solver SPECFEM3D-Cartesian (Komatitsch and Tromp, 2002) available at https://github.com/geodynamics/specfem3d. To generate the input data for the solver, the initial velocity model of the

435 Groningen gas field was constructed using gnam and PyAspect: available at https://github.com/code-cullison/gnam and https://github.com/code-cullison/pyaspect.

Data availability. The Groningen field subsurface models used in this study are available at https://nam-onderzoeksrapporten.data-app.nl/reports/download/groningen/en/3b4f8b0d-0277-40e0-8ff5-9a385c08327d (last accessed December 2021). The explanation of the models can be found in Romijn (2017).

440 *Author contributions.* LOMM conceptualized the methodology, performed the inversion, prepared the figures, and wrote the initial draft. TC helped in generating the elementary seismogram database and edited the manuscript. CW helped in developing the methodology and substantially improved the draft. All authors edited the manuscript.

Competing interests. All authors declare that they have no competing interests.

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